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Review Paper

Economic Growth and Environmental Pollution in Brunei: ARDL Bounds Testing Approach to Cointegration

ABSTRACT

Aims: This study examine the short run and long run dynamic relationship between economic growth and environmental pollution in Brunei. The study adopts the framework of Autoregressive Distributed Lag (ARDL) model to scrutinize the existence of the Environmental Kuznets Curve (EKC) among the studying variables, using time series data cover the period of 1974 to 2014.

Methodology: The ARDL bound test reveals the existence of long-run relationship among the integrated variables when CO₂ chosen as a dependent variable.

Results: The results support the existences of EKC hypotheses in the long-run whereas in the short-run an invested U-shaped curve was not confirmed between GDP and CO₂ in Brunei. The results of granger causality based on VECM analysis have shown unidirectional causality runs from economic growth to CO₂ in the short run. Further analysis through stability test indicates the coefficients in the model are stable and do not suffers with structural break within the time taken in the study.

Conclusion: The government of Brunei should proceeds to target the sustainable means of production which has environmental friendly and consumes less energy to enhance economic growth and maintain environmental quality in the long run.

Keywords: Carbon dioxide, Economic growth, ARDL, Granger causality, EKC, VECM, Brunei

1. INTRODUCTION

In recent years, many countries around the world have started to focus on an alternative means of production as the drastic actions to reduce alarming rate of environmental degradation. Similarly, it is well known that the poor productive capacities of a nation and severe climatic change have been the results of higher level of CO₂ emission (Abdoul & Hammami, 2017). The emission of carbon dioxide gas has grown severely due to various human activities, related to the expansion of land utilization as well as the rapid use of fossil fuels as a source of energy. Since the constant supply of energy is needed among the heavier industries in order to maintain the maximum production level, improving the human life and guarantee the strong economic growth among the nations (Salahuddin et al. 2015). Indeed, the rapid use of fossil fuel for production process has led the magnificent increases of CO₂ and other greenhouse gases in the atmosphere. These emitted substances not only destroy the natural environment but also bring the negative impacts to human life that are considered to be among the world's greatest environmental threats (Zahidul Islamet al.2013).

30 The economy of Brunei Darussalam mostly supported by the oil and gas industry which an
31 account for 60 percent of Gross Domestic Product (GDP) and 90 per cent of the total exports
32 in 2017 (OECD, ERIA, 2018, p. 206). "Brunei Darussalam is the fourth-largest oil producer
33 in South-East Asia and the ninth-largest exporter of liquefied natural gas (LNG) in the world"
34 (APEC, 2017). Although the Brunei Darussalam implemented some key actions directed to
35 reduce CO₂ emissions from fuel combustion that effectively begun in 2010. The overall
36 amount of carbon dioxide emissions had increased from 63.2 percent in 2010 to 67.5
37 percent in 2014, while the Methane gas (CH₄) had declined from 36.4 per cent in 2010 to
38 32.1percent in 2014 (UNFCCC, 2017). The CO₂ is the biggest contributor of Green House
39 Gases in the country which comes from the burning of fossil fuels that is widely used in the
40 electricity generation (48.7 per cent) and end-use sectors. Methane gas also has been
41 recorded to contribute significantly share to the total emission of GHGs, which generated
42 from the irregular releases of gas from industrial activities and land transportation accounts
43 for 12.6 percent and 34.5 percent respectively (UNFCCC, 2017). The strategies to reduce
44 the excessive uses of energy consumption in Brunei were due to the problem of increasing
45 energy efficiency. This might reduce the economic growth associated with collapse of
46 primary industries that depend mostly on combustion of fossil fuels like oil. Therefore, the
47 investigation of the relationship between energy consumption, urban population, economic
48 growth and CO₂ emissions is significant towards the implication of energy policies in Brunei.

49

50 **2. LITERATURE REVIEW**

51 This study tends to adopt the fundamental work of Simon Kuznets in 1955, whose study is
52 undertaken to investigate the inverse relationship between environmental quality and
53 economic growth (Kuznets, 1955). Environmental Kuznets Curve (EKC) hypothesizes that
54 there is a positive relationship between per capital income and environmental degradation
55 which is consistent up to the turning point, where the relationship overturned, that is to say,
56 at the initially stage, the environmental quality tends to decline as the economic growth rises
57 and eventually at the turning point it starts to improve with per capital [Saboori et al., 2012].
58 The environmental Kuznets curve draw the conclusion to either support the existence of
59 EKC hypotheses based on an inverted U-shaped or reject the hypotheses when results
60 portrays N-shaped EKC curve. The first empirical study on EKC hypotheses was conducted
61 by Grossman and Krueger in 1995 and 1993. In there an investigation employed specific
62 random city model and observed an Inverted U-shaped curve between various indicators of
63 environmental degradation (such as carbon dioxide emission) and GDP per capital
64 [Palamalai S. et al, 2015].

65 Moreover, there are numerous studies have been conducted to examine the linkages
66 between carbon dioxide emissions, economic growth and other controlled variables using
67 different econometric techniques to test the validity of EKC hypothesis and ended up with
68 strong evidences to either support or not the existence of EKC hypothesis. Balibey (2015)
69 examined the relationships between CO₂, economic growth and Foreign Direction
70 Investment (FDI) between 1974 and 2011, using 'VECM estimation' model. His results
71 verified the existence of EKC in Turkey. Furthermore, results of this study also
72 demonstrated the unidirectional granger causality runs from economic growth to
73 environmental degradation. Sambrano et al (2018) examine the existence of an inverted U-
74 shaped of EKC curve in Singapore using ARDL based on time series data over the period of
75 1971-2011. The empirical results support the EKC hypothesis both in long run and short run
76 phenomenon. Alabdulrazag and Alrajhi (2016) examine the relationship between economic
77 growth, CO₂, energy consumption and population density using ARDL bounds test to
78 cointegration and verify the validity of EKC hypothesis in KSA. Their results also support the
79 existence of inverted u-shaped in both short and long run.

80 In addition, Javaid and Zulfiqar (2017) investigate the links between CO₂, economic growth,
81 energy consumption, urban population and trade openness in Pakistan, based on time
82 series data between 1971 and 2010. Applying ARDL bound test to cointegration approach
83 and VECM to verify the existence of EKC curve in long run and short run phenomena
84 respectively. Their results support the Environmental Kuznets curve (EKC) between CO₂
85 and economic growth in long run as well as short run. In addition to that, other studies
86 revealed the similar results of supporting EKC hypotheses. [Osabuohien et al. 2014; Muftau
87 O et al. 2014; Xionglng T (2016); Zahidul Islam et al. 2017; Kebede, S 2017; Aye and Edoja,
88 2017; Bildiric M (2017); Naminse and Zhuang, 2017]

89 However, some other studies failed to admit the inverted U-shaped relationship with real-life
90 data in fact their results do not support EKC. Pandey and Mishra (2015) employ both
91 dynamic and static framework to examine the causality between CO₂ emissions and
92 economic growth using panel data cover the period between 1972 and 2010. The results of
93 this study failed to support the EKC hypothesis in SAARC countries. However the results
94 from VECM indicated the unidirectional granger causes from economic growth to carbon
95 dioxide. Linh and Lin (2014) examine the dynamic relationship between carbon dioxide
96 emission, economic growth, FDI and energy consumption between 1980 and 2010 using
97 granger causality and cointegration approach to verify the existence of EKC in Vietnam.
98 However their empirical results do not support EKC theory.

99 Furthermore, Muftau et al (2014) investigate the link between CO₂ and economic growth in
100 West Africa. Using fixed effect model to time series data cover the period between 1970 and
101 2011. The results indicate the N-shaped relationship between CO₂ and economic growth
102 which do not support the EKC theory in West Africa. Other previous studies failed to admit
103 EKC hypothesis Alam and Janifar, 2014; Ali A et al., 2015; Lacheheb et al (2015); Marsiglio
104 et al. 2016; and Jardón et al., 2017]. With respect to related reviews, it have been clearly
105 shown the contradiction in an existence of EKC hypotheses based on different techniques
106 run from individual to multilateral countries (cross sectional countries). To the best of our
107 knowledge, no study has conducted directly to examine the environment-growth nexus in the
108 context of Brunei using ARDL framework, Bound cointegration, Granger Causality test and
109 Sensitivity Analysis.

110

111 **3. METHODOLOGY**

112 **3.1 Theoretical Methodology**

113 Referring to the pioneering work of Grossman and Krueger in 1995, which is translated the
114 idea of EKC theory that explore the relationship between environmental degradation by
115 means of polynomial equation of per capita income. The standard estimation model can be
116 expressed as follows and it has been recently used by very other scholars include Saboori et
117 al (2012) and Lacheheb et al (2015).

$$118 \quad E = f(Y, Y^2, Y^3, Z) \quad (1)$$

119 Where E represents as environmental degradation as a function of Income (Y), Income
120 squared (Y²), cubic square (Y³) and a set of control variables (Z). In order to provide clear
121 interpretation of the coefficients, all variables converted into logarithm forms and plug into an
122 econometric model and therefore, the estimation model (1) will be:-

$$123 \quad \ln(CO_2) = \alpha_0 + \alpha_1 \ln Y_t + \alpha_2 (\ln Y_t)^2 + \alpha_3 (\ln Y_t)^3 + \alpha_4 (\ln Z_t) + \varepsilon_t \quad (2)$$

124 Where $\ln(\text{CO}_2)$ represents as environmental degradation as a proxy of carbon dioxide
 125 emission metrics tones per capital, Carbon dioxide has been used widely as explanatory
 126 variables (such as the study of Pandey and Mishar, 2015; Chen Lu, 2017; Ali et al, 2015;
 127 Jordan et al (2017); and Shaari et al (2017). Y as a GDP per capital, income squared (Y^2),
 128 income cubic (Y^3) are indicators of economic growth; and Z refers to others independent
 129 variables that may influence environmental pollution, ε stands as disturbance term, and t is a
 130 time series, α_s represent the elasticity of Y , Y^2 , Y^3 and Z respectively. This study included
 131 urban population and energy consumption as control variables. The inclusion of income (Y)
 132 and its exponential values as the determinants of environmental degradation into estimation
 133 model, may lead the model to suffer with perfect Multicollinearity (Usama *et al.*, 2016). In
 134 econometric theories, the presence of multicollinearity in the model leads to increases
 135 standard errors and further affects the hypothetical decision rules criterion. This
 136 circumstance has prompt researchers to test the possibility of presence of multicollinearity
 137 problem between income, income squares, income cubic and other control variables.

138 Table (01) indicates the results of correlation matrix among the proposed variables which
 139 intend to be included in the estimation model. The results in the table below depict the
 140 perfect correlation (strong correlation) between income ($\ln Y$), squared income ($\ln Y^2$) and
 141 cubic Income ($\ln Y^3$), and hence the decision to include them as key variables in the
 142 estimation model can creates the multicollinearity problem.

143 **Table 01: Correlation matrix**

	INCO2	INEC	INPOP	INY	INY2	INY3
INCO2	1.000000	.430549	.232410	-.364343	-.364641	-.364935
INEC	.430549	1.000000	.628180	-.472051	-.472591	-.473123
INPOP	.232410	.628180	1.000000	-.801483	-.802541	-.803587
INY	-.364343	-.472051	-.801483	1.000000	.999992	.999968
INY2	-0.364641	-0.472591	-0.802541	0.999992	1.000000	0.999992
INY3	-0.364935	-0.473123	-0.803587	0.999968	0.999992	1.000000

144
 145 Therefore our estimated model will drop all highly correlated variables and remained the rest
 146 of other explanatory variables. Therefore the model (2) will include $\ln \text{CO}_2$, $\ln \text{ENC}$, $\ln \text{POP}$
 147 and $\ln Y$ and expressed as follow:-

148
$$\ln(\text{CO}_2) = \alpha_0 + \alpha_1 \ln Y_t + \alpha_2 (\ln \text{ENC}_t) + \alpha_3 (\ln \text{POP}_t) + \varepsilon_t \quad (3)$$

149 Therefore, in order to avoid the inclusion of multicollinearity problem in the estimation model
 150 between income and its polynomial terms, Narayan and Narayan (2010) as cited in
 151 Alabdulrazag and Alrajhi (2016) suggest an alternative technique used to support or not the
 152 existence of EKC curve in developing economies. The decision compares the magnitude of
 153 elasticity of income with respect to CO_2 between long run and short run. If the results
 154 indicate that the long run effects on income elasticity is lower than the short run effects, this
 155 depicts that over a given period of time CO_2 falls as the income rises after turning point
 156 which will support the existence of EKC Hypotheses.

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160 **3.2 Data Sources**

161 This study employs the set of time series data collected from the World Bank Database
 162 (WDI) in 2018. The series of data collection that cover 1974-2014 include; carbon dioxide
 163 emission measured in metric tons per capita, Gross Domestic Product per capita as a
 164 constant \$ in 2010, energy consumption as kg of oil equivalent per capital and population as
 165 a percentage of urban population. All variables in this study are transformed into logarithms
 166 form namely as $\ln CO_2$, $\ln POP$, $\ln Y$ and $\ln ENC$.

Abbreviation	Variable	Measure	Previous Author on the same variable	Sources	Expected Sign
CO ₂	Carbon dioxide	Environmental pollution	J Alam, 2014	WDI	
Y	GDP	Economic growth	Omari, 2013	WDI	(+) short run
Y	GDP	Economic growth	Omari, 2013	WDI	(-)long run
ENC	Per capital energy use	Energy Consumption	Ghos et al, 2014	WDI	(+)
POP	Urban Population	Population	Javid and Zulfiqar, 2017	WDI	(+)

167 Author Compilation

169 **3.3 Estimation Methodology:**

170 **3.3.1 Unit Root Test**

171 Generally, the data collected from the large span of time period, the unit root problem would
 172 be possible to exist and without strictly careful, the results will be nonsense. In order to
 173 check the unit root problem in the data, the stationarity technique of time series should be
 174 applied such as "augmented dickey fuller" (ADF) (1979) and Phillips- Perron (PP) (1989)
 175 based on the following regression equation.

176
$$\Delta y_t = \alpha + \delta_1 + \beta y_{t-1} + \sum_{i=1}^k \gamma \Delta y_{t-1} + \mu_t \quad (4)$$

177 Where Δy_t indicates the first difference of y , μ_t represents the serial correlation errors and α ,
 178 δ , β and γ are parameters of the estimated model. If this test shows significant value, "it
 179 means that the variable series is stationary and does not has a unit root test, so the null
 180 hypothesis will be rejected but alternative hypothesis will be accepted, But if the stationary
 181 test is not significant, it means that the variable series is non-stationary and has a unit root
 182 test" (Shaari et al, 2012, p.102).

183 **3.3.2 Bound Testing Approach for Co-integration**

184 Bound testing approach for cointegration will be performed once after checking and
 185 confirming the stationarity of all variables that are integrated either in $I(0)$ or $I(1)$. This study
 186 employed 'Autoregressive distributed lag model (ARDL) technique' to check the stability of a
 187 long run and equilibrium relationship between CO_2 emission, economic growth, energy
 188 consumption and urban population. This technique is most powerful to cointegration since it
 189 generates more accurate results particularly for the case of small sample size (Javid and
 190 Zulfiqar, 2017).The Autoregressive Distribute Lag model can be estimated by using a simple

191 linear transformation technique to generate dynamic 'unrestricted error-correction model
 192 (UECM)'. This model tends to present the short run dynamism and an equilibrium position in
 193 the long run phenomenon. The following UECM is employed for equation 3 and expressed
 194 as follow

$$\begin{aligned}
 &195 \Delta \ln CO_2_t = \\
 &196 \alpha_0 + \alpha_{T^r} + \sum_{i=1}^p \varphi_i \Delta \ln CO_2_{t-1} + \sum_{i=0}^q \Omega_i \ln Y_t + \sum_{i=0}^r \vartheta_i \ln ENC_t + \sum_{i=0}^y \omega_i \ln FDI_t + \\
 &197 \lambda_C \ln CO_2_t + \lambda_Y \ln Y_{t-1} + \lambda_E \ln ENC_t + \lambda_P \ln POP_t + \mu_t \quad (5)
 \end{aligned}$$

198 Where $\varphi, \Omega, \vartheta$ and ω show the short run relationship in the given equation respectively and
 199 and the long run relationship among the coefficients variable is presenting by $\lambda_C, \lambda_Y, \lambda_E, \lambda_P$
 200 and λ_P . The null hypothesis of no co-integrated among the given variables is stated
 201 as $H_0: \lambda_C = \lambda_Y = \lambda_E = \lambda_P = 0$, while alternative hypothesis of co-integrated is given
 202 as $H_1: \lambda_C = \lambda_Y = \lambda_E = \lambda_P \neq 0$.

203 3.3.3 Error Correction Model Specification (ECM)

204 Once after confirming the existence of co-integration among the variables via bound test, the
 205 short and long run relation will be established by using VECM. It is possible to calculate the
 206 error correction term (ECT) from the long-run equation (Pesaran *et al*, 2001) by replacing the
 207 lagged level variables in the ARDL equation with ECT_{t-1} and estimate the model after impose
 208 the same optimal lags. The VECM model can be written as follows;

$$\begin{aligned}
 &209 \Delta \ln CO_2_t = \\
 &210 \alpha_0 + \sum_{i=1}^p \alpha_{1i} \Delta \ln CO_2_{t-1} + \sum_{i=1}^q \alpha_{2i} \ln Y_{t-1} + \sum_{i=1}^r \alpha_{3i} \ln ENC_{t-1} + \sum_{i=1}^y \alpha_{4i} \ln POP_{t-1} + \\
 &211 \lambda_1 ECT_{t-1} + \mu_t \quad (6)
 \end{aligned}$$

212 The $\ln CO_2$ is a function of its lagged values, lagged values of other exogenous variable in
 213 the model and the lagged value of ECT. The ECT_{t-1} is the one lagged error correction term
 214 indicates the co-integrating vectors and the speed of adjustments to equilibrium points
 215 presented by the coefficients of α_s . The presence of ECTs term in the model reveals that any
 216 change in the dependent variable as a result of the disequilibrium in the long run relationship
 217 and the changes in the independent variables. A negative and significant value of ECTs
 218 measures how much the error term is corrected itself each time towards the point of
 219 equilibrium in the long run

220 4.3.4 Granger Causality,

221 When the results of bound test supports the existence of long run relationship, there is an
 222 enough evidence to indicate that the underlying variables in Eq (3) are granger causes at
 223 least in one direction. Engle and Granger (1987) proposed that conducting the Granger
 224 causality test through Vector Auto Regressive (VAR) approach when long run relationship
 225 confirmed may provide inconsistent results in the presence of cointegration, therefore,
 226 adding of a new variable such as the Error Correction Term (ECT) to the VAR model would
 227 be helpful to explore the long run relationship (Alabdulrazag and Alrajhi, 2016, Shahbaz *et*
 228 *al.*, 2012). The direction of causality among the studying variable can be identify by the
 229 negative sign of the one lagged coefficient through ECT of the long run relations. The
 230 granger causality test through the framework of VECM techniques is expressed in the
 231 following equation:

$$(1-L) \begin{bmatrix} \Delta lco_t \\ \Delta ly_t \\ \Delta lENC_t \\ \Delta lPOP_t \end{bmatrix} = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \mu_3 \\ \mu_4 \end{bmatrix} + \sum_{t=1}^p (1_L) \begin{bmatrix} \pi_{11} & \pi_{12} & \pi_{13} & \pi_{14} \\ \pi_{21} & \pi_{22} & \pi_{23} & \pi_{24} \\ \pi_{31} & \pi_{32} & \pi_{33} & \pi_{34} \\ \pi_{41} & \pi_{42} & \pi_{43} & \pi_{44} \end{bmatrix} + \begin{bmatrix} \varphi \\ \Omega \\ \Phi \\ \omega \end{bmatrix} ECM_{t-1} + \begin{bmatrix} \eta_{1t} \\ \eta_{2t} \\ \eta_{3t} \\ \eta_{4t} \end{bmatrix} \quad (7)$$

233 Where the (1-L) indicates lag operator, 'ECM_{t-1} is lagged error-correction term'; Therefore,
 234 the main benefit of this model is to capture the causal relationship among the co-integrated
 235 parameters and to enables us to distinguish between short-run and long-run relations. The
 236 significance of the lagged error-correction term based on t-test(s) in the VECM, indicates the
 237 "long run" causal relationship, whereas the short run causal relationship is depicted through
 238 the significance of F-test of the lagged explanatory variables. After estimation of all
 239 equations, the next step is to perform diagnostic tests to validate the adequacy of the model.
 240 These diagnostic tests include serial correlation, normality distribution tests and
 241 Heteroscedasticity. The stability of the ECM performed using the cumulative sum (SUMUS)
 242 and cumulative sum squared (CUMUSQ) techniques to graphically confirm the stability of
 243 the variables estimated in the model (Pesaran et al., 2001 cited in Alabdulrazag and Alrajhi,
 244 2016).

245 4. Result Analysis and Discussion

246 4.1 Results of Unit Root Test

247 The bounds test framework is applicable for the variables that are either integrated in I (0) or
 248 I (1) (Pesaran et al, 2001). Therefore, the unit root tests are performed to determine the
 249 order of integration among the variables and to avoid any spurious results. The "Augmented
 250 Dickey-Fuller (1979) (ADF)" and "Phillips-Peron (1988) (PP)" unit root tests were performed
 251 to test the null hypotheses of unit root against the alternative hypotheses of stationarity. The
 252 results show all variables become stationary at I(0) or I(1) and not I(2). These results verify
 253 the adoption of ARDL model. The results of these tests are consistent with Ali et al (2015),
 254 Malibey M (2015) and Shaari (2017).

255 **Table 02 below presents the results of Unit Root Tests:**

Variable	Augmented Dickey-Fuller (ADF)		Phillips Perron (PP)	
	Constant Without Trend	Constant With Trend	Constant Without Trend	Constant With Trend
InCO2	-2.8173	-1.3443	-2.7931	-2.6275
InY	-.6251	-1.4312	-.3241	-1.6125
InENC	-2.4900	-3.5476*	-3.5475	-3.3981
InPOP	-2.1149	-1.1755	-9.2480**	-2.6797
First Difference I(1)				
ΔInCO2	-4.9363**	-5.3291**	-7.3386**	-6.6963**
ΔInY	-4.7898**	-4.7709**	-4.0141**	-4.1245**
ΔInENC	-5.8319**	*****	-12.5931**	-8.1187**
ΔInPOP	-3.1090*	-3.7755*	*****	-3.7755*

Note: ** and * denotes significant at 1%, and 5% significance level, respectively.

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259 **4.2 Bounds test approach to Co-integration**

260 The results of unit root test among the underlying variables indicate the application of the
 261 ARDL model to cointegration will give efficient and realistic estimates. The ARDL bounds
 262 test to cointegration investigate the possible existent of long run and short-run relationships
 263 among the underlying variables using the bound F-statistics in equation (5). Table (3)
 264 demonstrates the outcomes of 'ARDL bounds test to cointegration'. The results of this test
 265 confirm the existence of co integrated equation among the underlying variables, where the
 266 value of F-test is greater than the upper boundary of critical value at 5 percent. Therefore,
 267 we reject the null hypothesis of no co-integrated among the studying variables when InCO₂
 268 is dependent variables. Based on the results of cointegration using bound test, supports the
 269 estimation of VECM model of InCO₂ equation to determine the long and short run relation.

270 **Table 03 Results of bounds testing to cointegration**

Dep. Var.	Function	F-Statistics	C.V	Coint.	
InCO ₂	F _{InCO₂} (InCO ₂ / InY, InENC, InPOP)	5.0914	5%	Yes	ECM
InY	F _{InY} (InY/ InCO ₂ , InENC, InPOP)	1.7912	5%	No	ARDL
InENC	F _{InENC} (InENC/InCO ₂ , InY, InPOP)	3.1687	5%	No	ARDL
InPOP	F _{InPOP} (InPOP/InCO ₂ / InY, InENC,)	2.4405	5%	No	ARDL
Significant Value		Critical Values			
		Lower Bound I(0)		Upper Bound I(1)	
1%		2.72		3.77	
5%		3.23		4.35	
10%		4.29		5.61	

271

272 **4.3 Long-run estimation results**

273 The coefficient of economic growth is negative and statistical significant at 5 percent level as
 274 shown in table 04 below. The 1 percent increases in the economic growth in the long run is
 275 predicted the CO₂ emission to decreases by 5.33 percent. . This result is consistent with the
 276 findings of (Ahmed and Long, 2013) for Pakistan. The results also indicate one percent
 277 increases in the energy consumption holding other variables fixed the pollutants of CO₂
 278 emission will increases by 1.531 per cent in the long run. The coefficient of population
 279 growth also has significant impact to reduce CO₂ emission to more than 6 percent in the
 280 long run.

281 **Table 04: Long-run estimation results**

Long Run Coefficients				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
INEC	1.531756	.602970	2.540350*	.02
INPOP	-6.416853	2.321043	-2.764643**	.01
INY	-5.337163	2.063439	-2.586537*	.02
C	74.570173	27.958662	2.667158*	.02
Diagnostic test			Test-stats	p-value
Serial correlation			1.686965	.23

Normality			1.305386	.52
Heteroscedasticity			1.141277	.40

282 Note: ** and * denotes significant at 1%, and 5% significance level, respectively.

283

284 4.4 Short-run estimation

285 The results of the short-run relationship between carbon dioxide (CO₂), economic growth
286 (GDP), energy consumption (ENC) and population (POP) in Brunei are depicted in table (5).
287 The lagged value of the Error Correction model has a negative and statistical significant at 5
288 percent level. The coefficient of economic growth is negative and statistical significant. This
289 indicates that, economic growth in Brunei tends to improve the environmental quality in the
290 short run. This implies that a 1 percent increases in economic growth will leads to reduce the
291 level of CO₂ emission by 4.84 percent. The energy consumption has the positive effects on
292 the amount of carbon dioxide emissions. This depicts that a 1% increases in consumption of
293 energy will increases CO₂ emission by 0.89 percent. The results of population found to be
294 negative and significant at 5 percent level. This indicates that a percent rises in urban
295 population is predicted to reduce emitted of CO₂ gases by 52.38 percent in Brunei when
296 short run taken into consideration (Javaid and Zulfiqar, 2017). To examine the Inverted U-
297 shaped curve of Environmental Kuznets theory in Brunei, We compare the magnitude of
298 CO₂ with respect to economic growth between long run and short run. The estimation of
299 long-run and short-run as indicated in table 5 show that the negative elasticity of economic
300 growth on CO₂ emission in the short-run fail to support the existence of EKC whiles the
301 negative elasticity change in the long run tends to support the EKC hypotheses in Brunei.
302 This result supported with the study findings of [Ahmed and Long, 2013].

303 **Table 05: Short estimation results (lnCO₂)**

Co-integrating Form				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(INEC)	.888452	.263483	3.371953	.005
D(INEC(-1))	.168295	.297546	.565611	.58
D(INEC(-2))	-.399902	.295666	-1.352546	.20
D(INEC(-3))	-.363335	.249568	-1.455857	.17
D(INPOP)	-52.382543	299.7550	-.174751	.86
D(INY)	-4.843530	1.490686	-3.249194	.007
D(INY(-1))	-1.428382	2.900069	-.492534	.63
D(INY(-2))	4.050801	1.864779	2.172268	.05
CointEq(-1)	-1.143057	.309676	-3.691138	.003
Cointeq = INCO2 - (1.5318*INEC -6.4169*INPOP -5.3372*INY + 74.5702)				

304 Note: ARDL (2, 3, 1, 4) selected on the basis of AIC. ** and *** Represent 5% and 1%
305 level of significance, respectively. Dependent variable is lnCO₂.

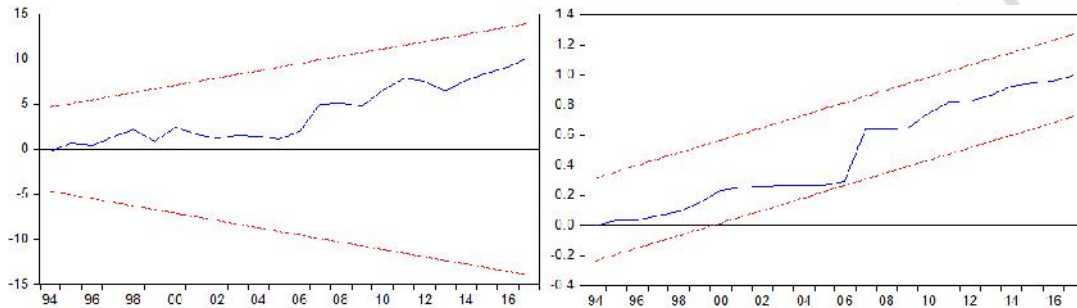
306

307 4.5 Stability Test and Sensitivity Analysis

308 The structural changes of macroeconomic policies in developing countries may likely to
309 causes the multiple structural breaks among the macroeconomic series. Therefore, it is vital
310 to verify the stability of long run and short run coefficients through 'the cumulative Sum
311 (CUSUM)'and 'Cumulative Sum of Squares (CUSUMSQ)' techniques which proposed by
312 Brown et al, 1975 as cited in Javaid and Zulfiqar, 2017. These tests do not require the

313 specification of dates where the structural breaks occurs, it only suggest that the parameters
 314 will be stable if the line passes within the bounds at 5 percent critical bounds. If the plot of
 315 these techniques passes outside the critical bound of 5 percent level of significance, reject
 316 the null hypothesis of not stable the regression coefficient and this implied that the coefficient
 317 in the ECM are not stable (Bekhet and Mater, 2013). The findings in the figure 1 below,
 318 indicate that the lines within the plots of 'CUSUM' and 'CUSUMSQ' test statistics passes
 319 within the critical bound of 5 percent for the studying period, and therefore, the coefficients
 320 are stable and do not suffer the structural change over the given time in the study.

321 **Figure 01: Stability Test**



322

323 **4.6 VECM Granger Causality Analysis**

324 The short-run co-integrated relations between CO₂, economic growth, energy consumption
 325 and population have shown the existence of causal relationship among the underlying
 326 variables. The short run causality has shown there is unidirectional significant causal
 327 relationship which runs from CO₂ to population. Similarly, there is also a unidirectional
 328 causal relationship which runs from economic growth (lnY) to carbon dioxide emission (CO₂)
 329 in the short run. The t-statistic of the ECT in the same table when carbon dioxide emission
 330 as a dependent variable reveals the existence of long run causal relationship which run from
 331 energy consumption, economic growth and population to CO₂ emission, Since the value of
 332 CO₂ emission is negative and statistical significance. There is no evidence of long run-
 333 relationships for DlnENC; DlnY and lnPOP, since, the soon after has a positive but
 334 insignificant Error correction term, while the two later have failed to pass the bounds of
 335 cointegration equation. The t-statistics of the ECT in the same table reveals the existence of
 336 long-run causal relationship among the mentioned variables. It is obvious to say that there is
 337 a bidirectional long-run causal relationship (equilibrium is corrected) between CO₂, lnY,
 338 INPOP and lnENC.

339 **Table 06: The results of Granger-Causality (Wald F-statistic test)**

	Short-run (Weak) Causality				Joint/Long-run (Strong) Causality				Long-run
	DICO ₂	DLY	DLENC	DLPOP	DICO ₂	DLY	DLENC	DLPOP	
DICO ₂	----	1.97	1.77	7.65*	----	5.53*	3.24#	4.47*	-.480*
DLY	3.62*	----	.08	0.14	----	----	----	----	-.017#
DLENC	2.92	3.32	----	1.12	----	----	----	----	.452*
DLPOP	2.05	.08	.09	----	----	----	----	----	-.001#

340 **Notes:** the null hypothesis is that there is no Granger causality between variables, *, **, #
 341 denote significant level at 1% and 5% and insignificant respectively.

342 **4. CONCLUSION**

343

344 This study examines the short run and long run dynamic relationships between CO₂, energy
345 consumption, urban population and economic growth in Brunei. It is one among the crucial
346 topic that deserves a special attention, since in most cases the economic growth in
347 developing economy associated with the environmental degradation. This happened due to
348 heavy consumption of pollutant substances as sources of energy in the industries and
349 means of transportation. The empirical results indicate the existence of long-run relationship
350 among the Co₂, energy consumption, urban population and economic growth. The results of
351 EKC found to be Inverted U-shaped curve indicates an evidence to support the EKC
352 hypothesis in Brunei. Therefore, the government of Brunei Darussalam should continue to
353 target the sustainable means of production which has environmental friendly and consumes
354 less energy to maintain the high rate of economic growth and improve the environmental
355 quality.

356 **AC**

357

358 **COMPETING INTERESTS**

359

360 "Authors have declared that no competing interests exist."

361

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