

# VECTOR AUTOREGRESSIVE MODELLING OF THE INTERACTION AMONG MACROECONOMIC STABILITY INDICATORS/VARIABLES IN NIGERIA (1981-2016)

## Abstract

Economic stability is a ~~major~~ essential macroeconomic goal for nations all over the world. This informed the desire by macroeconomic managers and investors alike for stable macroeconomic ~~conditions~~ proxy variables. However, the dynamic behaviour of ~~indicators of macroeconomic stability~~ these variables particularly their evolution, interaction and interdependence obviously cause shocks among themselves. This study therefore, is a multivariate time-series modelling and investigation of the interaction and pattern of causality among exchange rates, inflation rate, interest rates, and implicit price deflator in Nigeria using unrestricted Variance Autoregression (VAR). Quarterly data on the ~~study~~ variables from ~~spanning the period from~~ 1981 to 2016 were sourced from CBN Statistical bulletin and used for the study. Augmented Dickey Fuller Test (constant, and constant & linear trend) results showed that all variables were I(1) except interest rate I(0) and implicit price deflator I(2). The inverse root of AR characteristic polynomial showed that the VAR model was stable. The Trace Statistics and Max Eigen result showed no co-integrating relationship. The Schwarz Information Criterion showed a lag length of 2. The VAR estimates indicated that the exchange rate as well as inflation rates were significantly affected by their first lag, while exchange rate was significantly affected by its first and second lag. The system analysis particularly the Wald statistics showed that both lags of each variable were jointly significant in affecting itself. The impulse response showed that all variables were instantaneously affected by own shocks, however, it ruled out the response in exchange rate to contemporaneous shocks in inflation rate, interest rate and implicit price deflator. The variance decomposition further showed that at least 80% of the impulse response were from own shocks. It was consequently recommended that government should regulate these variables particularly interest rates and exchange rates while implicit price deflator and inflation rate should be stabilise. (ABSTRACT IS TOO LENGTHY. REDUCE IT TO 250 OR LESS!!!)

*Keywords: VAR, Impulse Response, Variance decomposition, Grangers Causality, Economic Stability*

## I. INTRODUCTION

Nigeria like other developing countries traditionally experienced macroeconomic instability. Conceptually, macroeconomic instability refers to a volatile macroeconomic condition. It is a phenomenon that makes the domestic macroeconomic environment less predictable. This is of concern because unpredictability hampers resource allocation decisions, investment, and growth.

Economic stability refers to absence of excessive fluctuation in key macroeconomic variables. An economy with fairly constant growth rate, low and fairly stable inflation, low and fairly stable interest rate, adequate and stable exchange rate. The World Bank describes a macroeconomic framework as stable "when the inflation rate is low and predictable, real interest rates are appropriate, the real exchange rate is competitive and predictable ... and the balance of payments situation is perceived as viable" (World Bank, 1990).

**Comment [h1]:** Why incorporate variables that are I(0) and I(2) in a VAR specification? Variables use for VAR modeling are not supposed to be of mixed order of integration

**Comment [h2]:** For variables that are of mixed order of integration, the Johansen test for cointegration (involving the Trace and Maximum Eigenvalue) CANNOT be applied to test for cointegration in time series modeling.

**Comment [h3]:** This is a hanging and incomplete statement

**Comment [h4]:** Add page number to this citation, please.

Economists obviously rely on multiple measures to achieve or guide stability, this paper analyses the maintenance or distortion in stability arising from the interaction among the identified stability variables using VAR approach.

**Comment [h5]:** Poorly written. Please recast and expand the introduction to reflect the motivation for the study, problem of the study, objective(s) of the study.

### Vector Autoregression (VAR)

Vector autoregression (VAR) is a technique used by macroeconomists to illustrate the joint dynamic behaviour of a collection of variables without requiring strong restrictions as required in the identification of fundamental structural parameters. VAR is an established method of time-series modelling; it has gained so much popularity since its introduction by Sims (1980).

VAR is a natural extension of the univariate autoregressive model; it depicts the dynamic behaviours of multivariate time series. The VAR model has proven to be very useful for financial time series, forecasting and describing the dynamic behaviour of economic time series. It often provides superior forecasts to models from univariate time series (Garba et al., 2017). Forecasts from VAR models are quite flexible because they can be made conditional on the potential future paths of specified variables in the model.

Although some useful applications of the estimates such as impulse-response functions (IRFs) or variance decompositions do require identifying restrictions, estimating the equations of a VAR does not require strong identification assumptions. Restrictions take the form of an assumption about the dynamic relationship between a pair of variables, for example, that exchange rate affect inflation rate only with a lag, or that exchange rate does not affect inflation rate in the long run.

A VAR system contains a set of  $m$  variables, each of which is expressed as a linear function of  $p$  lags of itself and of all of the other  $m - 1$  variables, including an error term.

VAR is a multivariate autoregressive linear time series model of the form

$$Y_t = \alpha + \sum_{i=1}^p \alpha_i Y_{t-i} + \varepsilon_t \quad (1)$$

Where;  $Y_t$  a set of  $n$  time series variables  $Y_t = (Y_{t1}, Y_{t2}, \dots, Y_{tn})$ , is a  $n \times 1$  Vector,  $\alpha_i$  are full rank  $m \times m$  matrix of coefficients, and  $i = 1, 2, 3, \dots, p$ ,

$U_t = (U_{t1}, U_{t2}, \dots, U_{tn})$  is an unobservable i.i.d. zero mean error term.

The reduced form of the unrestricted VAR model is a good approximation for the dynamic process of any vector of time series. This VAR estimation assumed a simple model for the stability variables of Nigerian economy with four endogenous variables: Exchange rate, Inflation rate, Interest rate, and implicit price deflator; the mathematical representation of a reduced form four variable

**Comment [h6]:** Wrongly placed. Please move this to Section III (Materials and Methods). Remember also to italicize 'et al' where it appears.

## II. LITERATURE REVIEW

Enders (1995), Lutkepohl (2001), Lutkepohl (2003) like other proponents of VAR suggest that in the forecasts of economic indicators, VAR models should be used as all variables in the models are endogenous, therefore, not a single variable may be removed when explanations for the behaviour of other variables are offered.

**Comment [h7]:** The literature review is too scanty. Please expand it by reviewing more relevant literature

Domac (2003) used VAR to study the relationship between the exchange rate, inflation, inflation expectations and money supply growth in 53 developing countries using annual data for the period from 1964-1998 to test the level-of-causal relationship between the aforementioned economic variables. The results from his work showed that 67% of the

**Comment [h8]:** Move this to Section III please.

variances in the rate of inflation in both long run and short run was explained by exchange rate depreciation and expected inflation explained about 10- 20% of movements in the rate of current inflation both in the short run and long run

Garba, Yahya, Babaita, Bankooko, and Amobi (2017) used VAR to model the structural relationships of exchange rates, of Naira to foreign currencies and concluded that Granger causality have been found useful in determining if one time series can be used in forecasting another, because it goes beyond correlation.

### III. MATERIALS AND METHODS

#### 3.1. Test for Stationarity

Time series data are often non stationary, however, the assumption of stationarity of the regressors and the regressand are crucial for the adoption of the Least Squares estimators (Etuck, 2012) in (Tuaneh& Essi, 2017). (Tuaneh, & Essi 2017) noted that the Stationarity of a series can strongly influence its behaviour, consequently, the use of non-stationary data can lead to spurious regression. Time series data on all variables included in the model are required to be stationary in order to carry out joint significant test on the lags of the variables. (Gujarati, 2013) explained that the various methods often used to test for stationarity; Augmented Dickey Fuller, the Philip Peron test, and the graphical method (the correlogram). The study however adopted the; Augmented Dickey Fuller Unit Root Test.

Augmented Dickey-Fuller (ADF) unit root test was employed to determine the order of integration of the series (i.e. to investigate the stationary status of each variable). The test is the  $t$ -statistic on the parameters. The following unit root tests regression equations are used for the first difference of the variables;

$$\Delta EXR_t = \tau_{11} + \tau_{12} \sum_{t-1}^k \rho_i \Delta EXR_{t-1} + \mu_{11} \quad (2)$$

$$\Delta IFR_t = \tau_{21} + \tau_{22} \sum_{t-1}^k \rho_i \Delta IFR_{t-1} + \mu_{12} \quad (3)$$

$$\Delta ITR_t = \tau_{31} + \tau_{32} \sum_{t-1}^k \rho_i \Delta ITR_{t-1} + \mu_{13} \quad (4)$$

$$\Delta IPD_t = \tau_{41} + \tau_{42} \sum_{t-1}^k \rho_i \Delta IPD_{t-1} + \mu_{11} \quad (5)$$

Where:  $\Delta$  is the difference operator

$U_t$  = random terms,  $t$  = time,  $k$  = number of lagged differences.

$\rho_i$  = coefficient of the preceding observation,  $(t-1)$  is the immediate prior observation,  $k$  is the number of lags, while  $\tau_{11}$ -  $\tau_{42}$  are the parameters to be determined.

The null hypothesis is that the series has a unit root 1(0), if ' $\tau$ ' is found to be more negative and statistically significant. We compare the  $t$ -statistic value of the parameter, with the critical value tabulated in (MacKinnon, 1991), We reject the null and conclude that the series do not have a unit root at levels

#### 3.2. Co-integration Test:

After examining the unit root of the study variables, and the order of integration of the series known, it is necessary to determine if there is a long run cointegrating relationship, since only

**Comment [h9]:** Please provide a description of the variables. What do EXR, IFR, ITR and IPD stand for?

variables that are of the same order of integration may constitute a potential cointegrating relationship.

Regression of one variable time series on one or more variables time series often can give spurious results; to guard against this is to find out if the series are cointegrated. Cointegration means despite being individually non-stationary, a linear combination of two or more time series can be stationary. This means subjecting these time series individually to unit root analysis and finding out if both are I (1) – non-stationary. Cointegration suggests that there is long-run or equilibrium relationship between them. To test whether the linear combination of the series that are non-stationary in levels are cointegrated (i.e. possesses a long-run equilibrium relationship). We employ the Johansen (1991), procedure of testing for a cointegrating relationship in a system of equations. The number of significant cointegrating vectors in nonstationary time series are tested by using the maximum likelihood based  $\lambda$ trace and  $\lambda$ max statistics introduced by Johansen and Juselius (1990). The stationary linear combination is called the cointegrating equation and interpreted as a long run relationship among the variables.

### 3.3. Models Specification-

Adapting equation (1) in the following VAR model form:

$$U(\text{VAR}) = (\text{EXR}, \text{INFL}, \text{INTR}, \text{IPD}) \quad (7)$$

With the lagged values of the endogenous variables and a constant being the exogenous variables, the VAR, may be written as:

$$\text{EXR}_t = \Gamma_{11(i)}\text{EXR}_{t-i} + \Gamma_{12(i)}\text{IFR}_{t-i} + \Gamma_{13(i)}\text{ITR}_{t-i} + \Gamma_{14(i)}\text{IPD}_{t-i} + \mathbf{K}_1 + \epsilon_{1t} \quad (8)$$

$$\text{IFR}_t = \Gamma_{21(i)}\text{EXR}_{t-i} + \Gamma_{22(i)}\text{IFR}_{t-i} + \Gamma_{23(i)}\text{ITR}_{t-i} + \Gamma_{24(i)}\text{IPD}_{t-i} + \mathbf{K}_2 + \epsilon_{2t} \quad (9)$$

$$\text{ITR}_t = \Gamma_{31(i)}\text{EXR}_{t-i} + \Gamma_{32(i)}\text{IFR}_{t-i} + \Gamma_{33(i)}\text{ITR}_{t-i} + \Gamma_{34(i)}\text{IPD}_{t-i} + \mathbf{K}_3 + \epsilon_{3t} \quad (10)$$

$$\text{IPD}_t = \Gamma_{41(i)}\text{EXR}_{t-i} + \Gamma_{42(i)}\text{IFR}_{t-i} + \Gamma_{43(i)}\text{ITR}_{t-i} + \Gamma_{44(i)}\text{IPD}_{t-i} + \mathbf{K}_4 + \epsilon_{4t} \quad (11)$$

One key feature of the equation is that no current time variables appear on the right-hand side of any of the equations. This makes it plausible, though not always certain, that the repressors are weakly exogenous.

However, equations (9) – (12) will be estimated if the variables are stationary at levels, in which case any shock to the stationary variables will be temporary. If the variables are nonstationary and not cointegrated, then they have to be transformed into stationary variables by differencing, if the variables are stationary after first difference and co-integrated then VAR can be transformed to vector error correction model (VECM).

### 3.4. VAR Lag Length Selection Criteria

The VAR lag length is selected using some model selection criteria. The general approach is to fit VAR models with orders  $p=0, 1, 2, \dots, p_{\max}$  and choose the value of  $p$  which minimizes the model selection criteria (Lutkepohl, 2005). Understanding that choosing too few lags could lead to systematic variation in the residuals whereas, too many lags come with the penalty of fewer degrees of freedom. The optimum or appropriate lag length for the VAR model was concluded based on the VAR lag order selection results in table 1, the researcher consequently concluded that the fit is good at lag 2 according to the Schwarz Information Criteria

**Table 1: VAR Lag Order Selection Results**

Lag	AIC	SC	HQ
0	39.69855	39.78421	39.73336
1	29.91182	30.77889	30.08589
2	30.00790	30.34015*	30.32121
3	29.54591	30.65958	29.99848
4	29.33480	30.79112	29.92661
5	28.94134*	30.74034	29.67241*
6	29.03831	31.17997	29.90863
7	29.10055	31.58487	30.11012
8	29.23601	32.06300	30.38482

\* indicates lag order selected by the criterion

The lag length selection criteria indicated two lags, hence the model above is written as

$$EXR_t = \Gamma_{111}EXR_{t-1} + \Gamma_{112}EXR_{t-2} + \Gamma_{121}IFR_{t-1} + \Gamma_{122}IFR_{t-2} + \Gamma_{131}INTR_{t-1} + \Gamma_{132}INTR_{t-2} + \Gamma_{141}IPD_{t-1} + \Gamma_{142}IPD_{t-2} + K_1 + \varepsilon_{1t} \quad (12)$$

$$IFR_t = \Gamma_{211}EXR_{t-1} + \Gamma_{212}EXR_{t-2} + \Gamma_{221}IFR_{t-1} + \Gamma_{222}IFR_{t-2} + \Gamma_{231}INTR_{t-1} + \Gamma_{232}INTR_{t-2} + \Gamma_{241}IPD_{t-1} + \Gamma_{242}IPD_{t-2} + K_2 + \varepsilon_{2t} \quad (13)$$

$$ITR_t = \Gamma_{311}EXR_{t-1} + \Gamma_{312}EXR_{t-2} + \Gamma_{321}IFR_{t-1} + \Gamma_{322}IFR_{t-2} + \Gamma_{331}INTR_{t-1} + \Gamma_{332}INTR_{t-2} + \Gamma_{341}IPD_{t-1} + \Gamma_{342}IPD_{t-2} + K_3 + \varepsilon_{3t} \quad (14)$$

$$IPD_t = \Gamma_{411}EXR_{t-1} + \Gamma_{412}EXR_{t-2} + \Gamma_{421}IFR_{t-1} + \Gamma_{422}IFR_{t-2} + \Gamma_{431}INTR_{t-1} + \Gamma_{432}INTR_{t-2} + \Gamma_{441}IPD_{t-1} + \Gamma_{442}IPD_{t-2} + K_4 + \varepsilon_{4t} \quad (15)$$

The researcher used Eviews 8 in the statistical data analysis which requires a different model specification, for the purpose of analysis in the Eviews, the model is specified as:

VAR Model Specification (Eviews): LS 1 2 EXR IFR ITR YT @ C

$$EXR = C(1,1)*EXR(-1) + C(1,2)*EXR(-2) + C(1,3)*IFR(-1) + C(1,4)*IFR(-2) + C(1,5)*ITR(-1) + C(1,6)*ITR(-2) + C(1,7)*IPD(-1) + C(1,8)*IPD(-2) + C(1,9) \quad (16)$$

$$IFR = C(2,1)*EXR(-1) + C(2,2)*EXR(-2) + C(2,3)*IFR(-1) + C(2,4)*IFR(-2) + C(2,5)*ITR(-1) + C(2,6)*ITR(-2) + C(2,7)*IPD(-1) + C(2,8)*IPD(-2) + C(2,9) \quad (17)$$

$$ITR = C(3,1)*EXR(-1) + C(3,2)*EXR(-2) + C(3,3)*IFR(-1) + C(3,4)*IFR(-2) + C(3,5)*ITR(-1) + C(3,6)*ITR(-2) + C(3,7)*IPD(-1) + C(3,8)*IPD(-2) + C(3,9) \quad (18)$$

$$IPD = C(4,1)*EXR(-1) + C(4,2)*EXR(-2) + C(4,3)*IFR(-1) + C(4,4)*IFR(-2) + C(4,5)*ITR(-1) + C(4,6)*ITR(-2) + C(4,7)*IPD(-1) + C(4,8)*IPD(-2) + C(4,9) \quad (19)$$

The system of equation above can also be presented in Eviews for ease of analysis, explanation and understanding as:

$$EXR = C(1) * EXR(-1) + C(2) * EXR(-2) + C(3) * IFR(-1) + C(4) * IFR(-2) + C(5) * ITR(-1) + C(6) * ITR(-2) + C(7) * IPD(-1) + C(8) * IPD(-2) + C(9) \quad (20)$$

$$\text{IFR} = \text{C}(10) * \text{EXR}(-1) + \text{C}(11) * \text{EXR}(-2) + \text{C}(12) * \text{IFR}(-1) + \text{C}(13) * \text{IFR}(-2) + \text{C}(14) * \text{ITR}(-1) + \text{C}(15) * \text{ITR}(-2) + \text{C}(16) * \text{IPD}(-1) + \text{C}(17) * \text{IPD}(-2) + \text{C}(18) \quad (21)$$

$$\text{ITR} = \text{C}(19) * \text{EXR}(-1) + \text{C}(20) * \text{EXR}(-2) + \text{C}(21) * \text{IFR}(-1) + \text{C}(22) * \text{IFR}(-2) + \text{C}(23) * \text{ITR}(-1) + \text{C}(24) * \text{ITR}(-2) + \text{C}(25) * \text{IPD}(-1) + \text{C}(26) * \text{IPD}(-2) + \text{C}(27) \quad (22)$$

$$\text{IPD} = \text{C}(28) * \text{EXR}(-1) + \text{C}(29) * \text{EXR}(-2) + \text{C}(30) * \text{IFR}(-1) + \text{C}(31) * \text{IFR}(-2) + \text{C}(32) * \text{ITR}(-1) + \text{C}(33) * \text{ITR}(-2) + \text{C}(34) * \text{IPD}(-1) + \text{C}(35) * \text{IPD}(-2) + \text{C}(36) \quad (23)$$

This is an indication that 36 parameters would be estimated. The square of the number of variables multiplied by the number of lags plus the number of variables  $[(4^2)2 + 4] = 36$

## IV. RESULTS

### 4.1 Time Plots

The time plots shown in figure 1 to figure 4 are indications that all variables showed fluctuations within the period of the study, no variable followed a steady trend.

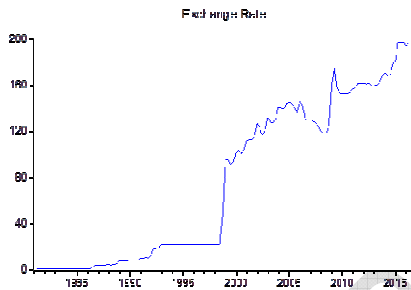


Figure 1: Time plot of Exchange Rate

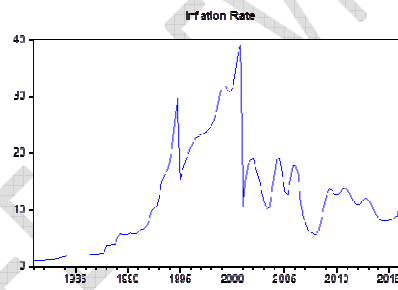


Figure 2: Time plot of Inflation Rate

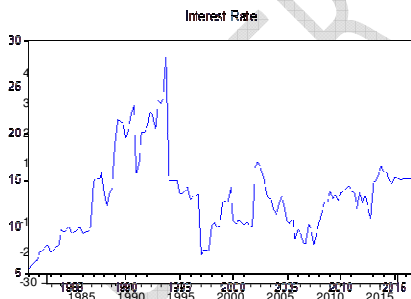


Figure 3: Time plot of Interest Rate

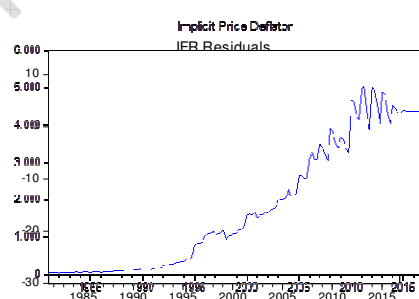
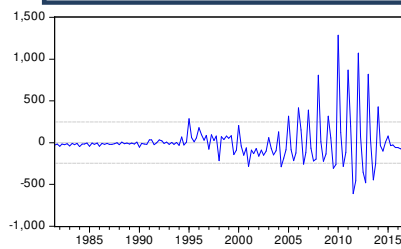
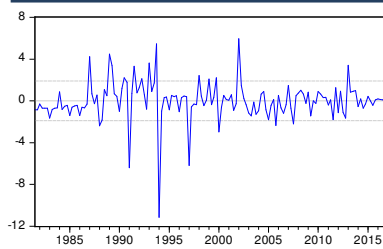


Figure 4: Time plot of Implicit Price



**Figure 5: Residual Plots at levels on all Variables**

**Comment [h10]:** Where is the figure, please?

## 4.2. Diagnostic Test Results

### 4.2.1. Unit Root Test Result

Since the study variables involved time series data, the Johansen technique cannot be applied unless it is established that the variables concerned are stationary. Data on each series were tested for stationarity so as to avoid the problem of spurious regression (Tuaneh and Essi, 2017). For this study, the Augmented Dickey-Fuller (ADF) test was used to test the null hypothesis of a unit root. The null hypothesis of a unit root is rejected in favour of the stationary alternative in each case if the test statistic is more negative than the critical value. A rejection of the null hypothesis means that the series do not have a unit root.

Table 2 presents results of the unit root tests, p-values are in brackets. The results showed that at levels, all variables had unit root (p-values > 0.05), however, all variables do not have unit root at levels (t-values more negative than the test statistics at 99% confidence, more so p-values are less than 0.05 level of significance at both intercept, and Constant & trend, consequently the null hypothesis of unit roots were rejected. Conclusively, Exchange rate, Inflation Rate, Interest Rate and Implicit price deflator were stationary at order 1(1).

**Comment [h11]:** Compare this with what you reported in the abstract

**Table 2: Augmented Dickey-Fuller Unit Root Test Result**

Variables	Levels		1st Difference		Order of Integration
	Constant	Constant, Linear Trend	Constant	Constant, Linear Trend	
Exchange Rate (EXR <sub>t</sub> )	0.0538(0.96)	-2.3907(0.38)	-6.6041 (0.000)	-6.6435 (0.000)	1(1)
Inflation Rate (IFR <sub>t</sub> )	-2.2331(0.19)	-2.2931(0.43)	-9.6703 (0.000)	-9.6426 (0.000)	1(1)
Interest Rate (ITR <sub>t</sub> )	-3.0371(0.03)	-2.9982(0.13)	-9.9293 (0.000)	-9.9090 (0.000)	1(1)
Implicit Price Deflator (IPD <sub>t</sub> )	0.0942(0.96)	-2.2261(0.47)	-4.8860 (0.000)	-4.9357 (0.000)	1(1)
<b>Test critical values:</b>	<b>%level</b>	<b>Constant</b>	<b>Constant, Linear Trend</b>		
	1% level	-3.4768	-4.0239		
	5% level	-2.8818	-3.4417		
	10% level	-2.5776	-3.1454		

**Comment [h12]:** This result is different from what is reported in the abstract. Please revisit the abstract and effect necessary corrections

### 4.2.2. Co-integration Test Result

The long run combination of stationary processes can be non stationarity. Cointegration exists if two variables have a long run equilibrium relationship between them. This study employs the Johansen maximum likelihood approach to test for co-integration. Though trace statistic is said to be more robust to both skewness and excess kurtosis in residuals than the maximum-eigen value test, the Johansen maximum likelihood approach is said to be more preferable to the other methods due to its properties (Wassell and Saunders, 2000) the researcher consequently used both maximum-eigen test and the trace statistics.

Table 3 showed the result of the  $\lambda_{\text{trace}}$  and  $\lambda_{\text{max}}$  statistics respectively. Max-eigenvalue test and Trace test indicates no co-integration at the 0.05 level

**Table 3: Johansen Co-integration Test Result**

Hypothesized No. of CE(s)	Eigenvalue	Trace			Max-Eigen		
		Statistic	0.05 Critical Value	P**	Statistic	0.05 Critical Value	P**
None	0.1225	38.0860	47.8561	0.2983	18.0379	27.5843	0.4920
At most 1	0.0912	20.0481	29.7970	0.4196	13.2054	21.131	0.4335
At most 2	0.0480	6.8427	15.4947	0.5960	6.7938	14.264	0.5138
At most 3	0.0003	0.0489	3.8414	0.8249	0.0489	3.8414	0.8249

Max-eigenvalue test and Trace test indicates no co-integration at the 0.05 level

**4.3. VAR Analysis Result of the Contemporaneous Coefficients**

$$\text{EXR}_t = 1.25\text{EXR}_{t-1} - 0.29\text{EXR}_{t-2} + 0.26\text{IFR}_{t-1} - 0.15\text{IFR}_{t-2} - 0.10\text{ITR}_{t-1} - 0.07\text{ITR}_{t-2} - 0.001\text{IPD}_{t-1} + 0.002\text{IPD}_{t-2} + 0.35$$

$$\text{IFR}_t = 0.03\text{EXR}_{t-1} - 0.03\text{EXR}_{t-2} + 0.90\text{IFR}_{t-1} + 0.02\text{IFR}_{t-2} + 0.06\text{ITR}_{t-1} + 0.12\text{ITR}_{t-2} - 0.00007\text{IPD}_{t-1} + 0.0003\text{IPD}_{t-2} + 0.09$$

$$\text{ITR}_t = 0.02\text{EXR}_{t-1} - 0.02\text{EXR}_{t-2} + 0.005\text{IFR}_{t-1} - 0.03\text{IFR}_{t-2} + 0.80\text{ITR}_{t-1} + 0.91\text{ITR}_{t-2} - 0.0002\text{IPD}_{t-1} - 0.0002\text{IPD}_{t-2} + 11.51$$

$$\text{IPD}_t = -1.09\text{EXR}_{t-1} + 3.60\text{EXR}_{t-2} + 0.50\text{IFR}_{t-1} - 1.17\text{IFR}_{t-2} - 2.12\text{ITR}_{t-1} + 3.99\text{ITR}_{t-2} + 0.82\text{IPD}_{t-1} + 0.06\text{IPD}_{t-2} + 39.94$$

The estimated model (substituted coefficients) above is a representation of the detail VAR estimation output. The estimates of the coefficients of multiple determinations ( $R^2$ ) of the models were respectively 0.992, 0.883, 0.808, and 0.979 respectively indicating that the dependent variables were largely explained by the independent variables. The Durbin Watson statistics were 1.82, 2.03, 2.03, and 2.12 respectively, hence there was no reason to suspect serial autocorrelation. The VAR estimates indicate that exchange rate, inflation rates, interest rates, and implicit price deflator were positively and significantly affected by their own first lags. Only exchange rate was significantly affected by its own second lag. The system analysis particularly the Wald statistics showed that both lags of each variable were jointly significant in affecting itself.

The VAR result above satisfy the stability condition as no root lies outside the unit root circle as shown in graph of the inverse roots of a characteristic polynomial in figure 6 below. More so, the table 4 showed that the modulus less than one but greater than zero

**Comment [h13]:** The R-squared and the Durbin Watson statistic are not reported anywhere in the paper.



**Table 4: Roots of Characteristic Polynomial (Endogenous variables: EXR IFR ITR IPD, Exogenous variables: C)**

Root	Modulus
0.997994	0.997994
0.919457 - 0.043599i	0.920490
0.919457 + 0.043599i	0.920490
0.836571	0.836571
0.357837	0.357837
-0.091403	0.091403
-0.074191 - 0.049922i	0.089423
-0.074191 + 0.049922i	0.089423

VAR satisfies the stability condition.

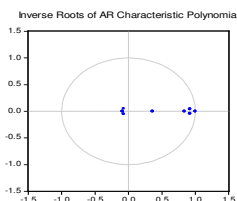


Figure 6: Inverse roots of a Characteristic Polynomial

No root lies outside the unit circle.

#### 4.4. Granger Causality

The granger causality test conducted and the summary result presented in table 5 below showed that only the combine lags (lag 1 and lag 2) of exchange rates granger caused implicit price deflator ( $PV = 0.022 < 0,05$ ). Inflation rates (lag 1 and lag 2) taken together do not granger cause exchange rates, interest rates and implicit price deflator taken diagonally from top to bottom. Similarly, the lags of interest rates jointly do not granger cause exchange rates, inflation rates and implicit price deflator. The probability values in the last column of table 3 indicate that the lags of all the independent variables taken together do not granger cause the dependent variables.

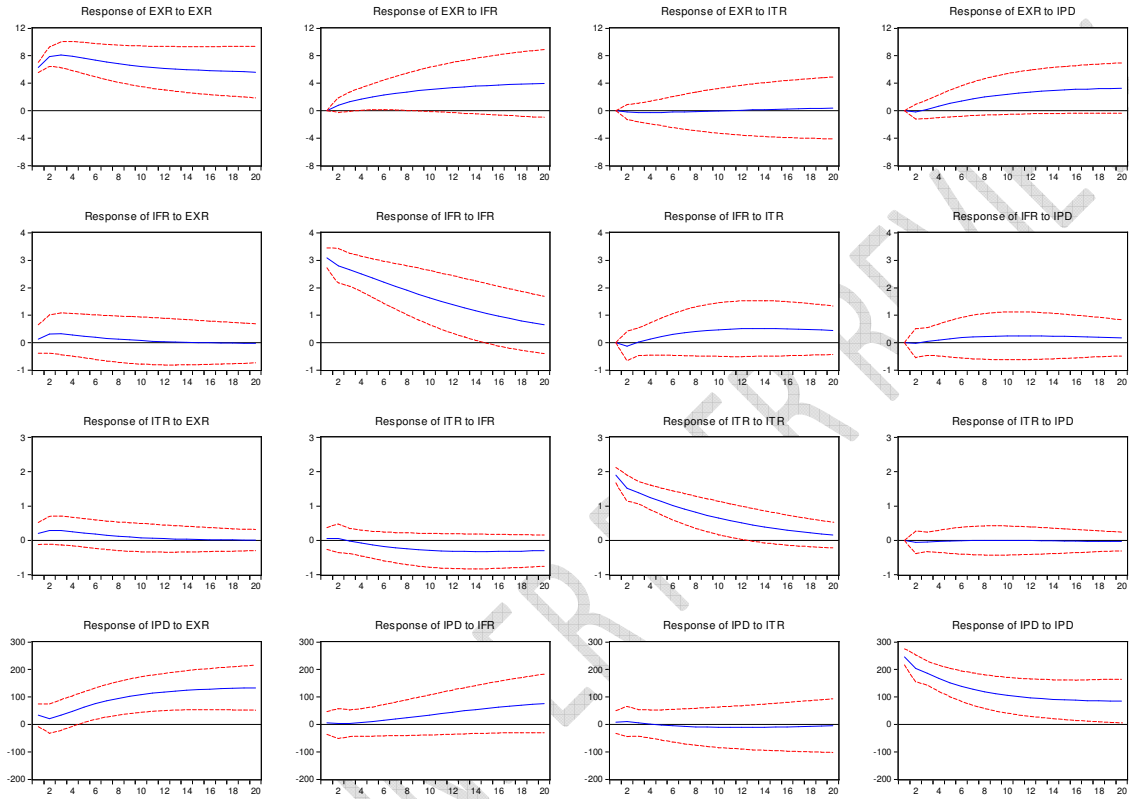
Most notably, the combine lags of each variable significantly affected itself. Exchange rates (lag 1 and lag 2) significantly caused current exchange rate ( $\chi^2 = 1755.4, P = 0.000$ ). Inflation rates (lag 1 and lag 2) significantly caused current Inflation rates ( $\chi^2 = 862.1, P = 0.000$ ). Interest rates (lag 1 and lag 2) significantly caused current Interest rates ( $\chi^2 = 546.3, P = 0.000$ ). Implicit Price Deflator (lag 1 and lag 2) significantly caused current Implicit Price Deflator. ( $\chi^2 = 583.2, P = 0.000$ )

**Table 5: Granger Causality (Block Exogeneity Wald and System Wald) Test Result (Test Statistics is Chi-square and P-values in Bracket)**

Dependent Variables	Independent Variables				
	EXR	IFR	ITR	IPD	All
Exchange Rate (EXR <sub>t</sub> )	<b>1755.4(0.00)*</b>	4.524(0.10)	0.142(0.93)	3.019 (0.22)	6.903(0.32)
Inflation Rate (IFR <sub>t</sub> )	1.044(0.59)	<b>862.1(0,00)*</b>	1.621(0.44)	0.277(0.87)	2.537(0.86)
Interest Rate (ITR <sub>t</sub> )	0.760(0.68)	2.095(0.35)	<b>546.3(0.00)*</b>	0.123(0.94)	2.883(0.82)
Implicit Price deflator (IPD <sub>t</sub> )	<b>7.566(0.02)*</b>	0.081(0.95)	0.222(0.89)	<b>583.2(0.00)*</b>	8.733(0.18)

#### 4.6. Impulse Response

Response to Cholesky One S.D. Innovations  $\pm 2$  S.E.



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### Figure 7: Impulse Response graphs

The zero values right from the start at lag zero for the immediate or contemporaneous response is to shocks are imposed by the Cholesky decomposition by the particular ordering. The first column of figure 7 represents response of exchange rates to shocks to all other variables, the second column represents variations in inflation rates to shocks to all other variables, the third column represents changes in interest rates to shocks to all other variables, while the fourth column represents response of interest rates to shocks to all other variables.

#### 4.6.1. Impulse Response of Exchange Rates

The first row of figure 7 above shows the response of exchange rates to shocks to exchange rates, inflation rates, interest rates and implicit price deflator. The zero values right from the start at lag zero ruled out to have an immediate effect. Consequently, exchange rate had an immediate and positive response to shocks in exchange rates, it however did not have an immediate nor positive response to shocks in inflation rates, interest rates and implicit price deflator, the response to interest rates is not immediate nor subsequently.

**Comment [h14]:** Not comprehensive, not detailed enough.

#### 4.6.2. Impulse Response of Inflation Rates

The second row of figure 7 above shows the response of inflation rates to shocks to in all studied variables. Inflation rates had an immediate and positive response to shocks in inflation rates, it however did not have an immediate response to shocks in exchange rates, interest rates and implicit price deflator, the response to exchange rates and implicit price deflator were not immediate nor subsequently.

**Comment [h15]:** Poorly explained. Not comprehensive, not detailed enough

#### 4.6.3. Impulse Response of Interest Rates

Row 3 of figure 7 shows the response of interest rates to shocks to all variables of the study. Interest rates had an immediate and positive response to shocks in interest rates, it however did not have an immediate response to shocks in exchange rates, inflation rates and implicit price deflator, the response to exchange rates, inflation rates, and implicit price deflator were not immediate nor subsequently.

**Comment [h16]:** Poorly discussed. Not comprehensive, not detailed enough

#### 4.6.3. Impulse Response of Implicit Price Deflator

Row 4 of figure 7 shows the response of implicit price deflator to shocks to all variables of the study. implicit price deflator had an immediate and positive response to shocks in itself and exchange rates, it however did not have an immediate response to shocks in exchange rates, inflation rates and interest rate, the response to inflation rates was not immediate nor subsequently.

**Comment [h17]:** Not detailed enough. Please elaborate.

### 4.7. Variance Decomposition

#### 4.7.1 Variance Decomposition of Exchange Rates

The first section of table 6 shows that in the short run, the response of exchange rate due to own shock is 98.5%. The table also showed that a shock in inflation rates, interest rate and implicit price deflator can respectively cause 1.3%, 0.06%, and 0.03% fluctuations in exchange rates. In the long run however, the response of exchange rate due to own shock is

41 88.53%. The fluctuations in exchange rates due to impulse in inflation rates, interest rate and  
 42 implicit price deflator are 7.82%, 0.06%, and 3.57% respectively.

**Comment [h18]:** Please consult appropriate texts and/or relevant journal articles and do a comprehensive and accurate discussion of the results. For example, 7.82% of short run intertemporal variation/fluctuation in exchange rate is due to exogenous shock to inflation in the 10<sup>th</sup> year...

43 **4.7.2 Variance Decomposition of Inflation Rates**

44 The responses of inflation rates in the short run due to own shock as indicated in the second  
 45 section of table 6 shows is 97.25%. The shock in exchange rates, interest rate and implicit  
 46 price deflator can respectively cause 0.88%, 0.06%, and 0.008% fluctuations in inflation  
 47 rates. In the long run however, the response of inflation rate due to own shock is 97.15%. The  
 48 fluctuations in inflation rates due to impulse in exchange rates, interest rate and implicit price  
 49 deflator are 0.79%, 1.57%, and 0.47% respectively.

**Comment [h19]:** Not well explained. Please consult appropriate texts and do a comprehensive and accurate discussion of the results.

50 **4.7.3 Variance Decomposition of Interest Rates**

51 The responses of interest rates in the short run due to own shock as indicated in the third  
 52 section of table 6 shows is 99.03%. The shock in exchange rates, inflation rates and implicit  
 53 price deflator can respectively cause 2.59%, 0.08%, and 0.06% fluctuations in interest rates.  
 54 In the long run however, the response of interest rate due to own shock is 95.05%. The  
 55 fluctuations in interest rates due to impulse in exchange rates, inflation rates, and implicit  
 56 price deflator are 0.79%, 1.57%, and 0.47% respectively.

**Comment [h20]:** Please consult appropriate texts and do a comprehensive and accurate discussion of the results.

57 **4.7.4 Variance Decomposition of Implicit Price Deflator**

58 The fluctuations in implicit price deflator in the short run due to own shock as shown in the  
 59 third section of table 6 shows is 97.903%. The shocks in exchange rates, inflation rates and  
 60 interest rates can respectively cause 2.59%, 0.08%, and 0.06% fluctuations in implicit price  
 61 deflator. However, in the long run, the response of implicit price deflator due to own shock is  
 62 82.07%. The fluctuations in implicit price deflator due to impulse in exchange rates, inflation  
 63 rates, and interest rates are 0.79%, 1.57%, and 0.47% respectively.

**Comment [h21]:** Please consult appropriate texts and do a comprehensive and accurate discussion of the results.

64 **Table 6: Variance Decomposition Result**

Period	S.E.	EXR	IFR	ITR	IPD
<b>Variance Decomposition of EXR:</b>					
3	13.04274	98.51092 (1.66923)	1.381123 (1.46291)	0.069878 (0.68870)	0.038083 (0.70566)
10	24.34593	88.53746 (9.25032)	7.827635 (6.67171)	0.060508 (3.30919)	3.574401 (4.86566)
<b>Variance Decomposition of IFR:</b>					
3	4.969223	0.886573 (2.72610)	99.03950 (3.04236)	0.065738 (0.95249)	0.008185 (0.76755)
10	7.506602	0.794994 (4.53409)	97.15740 (6.94553)	1.570381 (4.40059)	0.477221 (2.78445)
<b>Variance Decomposition of ITR:</b>					
3	2.837308	2.595769 (4.13352)	0.084582 (1.66767)	97.25390 (4.84294)	0.065752 (0.80308)
103	3.859087	2.689754 (5.09400)	2.203673 (4.40458)	95.05995 (7.21623)	0.046620 (2.84721)
<b>Variance Decomposition of IPD:</b>					
3	373.7500	1.913927 (3.23674)	0.037748 (1.35559)	0.145731 (1.69396)	97.90259 (3.91306)
10	564.6483	16.69690 (12.2860)	1.095751 (3.95528)	0.179325 (4.12169)	82.02803 (12.2815)

67 **V. CONCLUSION**

68 During the period considered, the combined lags of exchange rates, inflation rates, interest  
 69 rates, and implicit price significantly caused own shocks, however, fluctuations due to other  
 70 study variables were minimal as shown by the impulse response and variance decomposition  
 71 analyses. Worthy of note is that; the study ruled out the response of exchange rate to  
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73 contemporaneous shocks in inflation rate, interest rate and implicit price deflator, it also rule  
74 out the fluctuation of inflation rate to contemporaneous impulse in exchange rate, interest rate  
75 and implicit price deflator and finally ruled out the response of interest rate to  
76 contemporaneous shocks in inflation rate, exchange rate and implicit price deflator. The test  
77 of significance particularly the granger causality test indicated significant influence of a  
78 particular variable by its combine lags. More-so, the causality between exchange rates and  
79 implicit price deflator was significant and uni-direction from exchange rates to implicit price.

**Comment [h22]:** Do you mean implicit price deflator?

80 Since own shocks have been found to be major and significant determinants of impulse, it is  
81 recommends that economic modelling should include lags of the dependent variable as  
82 independents, particularly for multivariate models. It is also recommended that government  
83 regulates this variables particularly interest rates and exchange rates while implicit price  
84 deflator and inflation rate should stabilise

**Comment [h23]:** Do you mean 'these'?

**Comment [h24]:** Stabilise what?

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## 86 REFERENCES

**Comment [h25]:** Some authors cited in-text are not listed here. Please correct this. For example, Johansen (1991), Johansen and Juselius (1990) is not listed here

87 Cuvak, A., and Kalinauskas, Z. (2009). Application of vector autoregression model for  
88 Lithuanian inflation. *Ekonomika IR Vadyba: Economics and Management* (14), 145-  
89 150

90 Domac, I. (2003). Explaining and forecasting inflation in Turkey. *Turkey Central Bank  
91 Publication*

92 Enders, W. (1995). *Applied Econometric Time Series*. New York: John Wiley & Sons.

93 Etuk, E. H. (2012). Predicting inflation rates of Nigeria using a seasonal Box-Jenkins model,  
94 *Journal of Statistical and Econometric Methods*, 1(3)(page numbers???)

95 Garba, M, K., Yahya, W. B., Babaita, H. T., Bankooko, A. W., and Amobi, A. Q. (2017).  
96 Modeling structural relationships of exchange rates, of Naira to foreign currencies.  
97 *Nigerian Statistical Society*. 1, 41-47

**Comment [h26]:** Please check this up and correct it. This cannot be the name of a journal

98 Gujarati, D. N. (2013). *Basic Econometrics*. McGraw Hills:Glasgow.

99 Lutkepol, H. (2001). Vector Autoregressive and Vector Error Correction Models. Institute of  
100 Statistics and Econometrics.

**Comment [h27]:** Poor referencing

101 Lutkepol, H. (2003). Vector Autoregressive Models. Companion to theoretical econometrics.

**Comment [h28]:** Poor referencing

102 Sims, C. (1980). Macroeconomics and reality. *The Econometrics*, (48), 1-48.

103 Tuaneh, G. L., and Essi, I. D. (2017) Simultaneous equation modelling the investment and  
104 consumption function in Nigeria. *International Journal of Economics and Business  
105 Management* 3(8), 53-71

106 World Bank (1990). *Adjustment Lending Policies for Sustainable Growth. Policy and  
107 Research*. Series no. 14. Oxford University Press, Washington DC.

108